

# Limit behaviour of random walks with elastic screens

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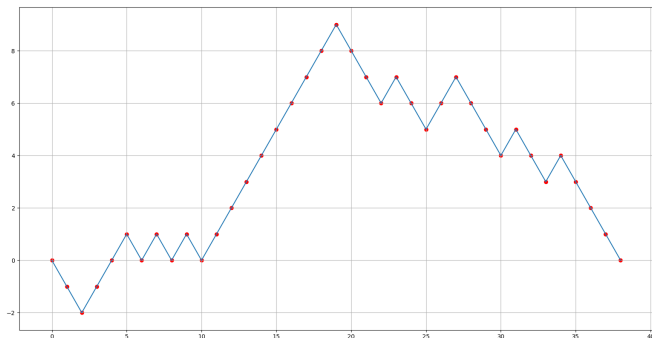
Let  $\{S(n)\}$  be a simple random walk, i.e.

$$S(n) = \sum_{i=1}^n \xi_i,$$

where  $\{\xi_i\}$  are independent  $\pm 1$  with probability  $\frac{1}{2}$ .

Extend  $S(n)$  for all  $t$  by linearity:

$$S(t) = S(n) + (t - n)(S(n + 1) - S(n)), \quad t \in [n, n + 1].$$



## Problem

Consider a random walk  $\{\tilde{S}(n)\}$ , which behaves as  $\{S(n)\}$  everywhere except 0. In 0 it stops for some random amount of time and then continues its way as  $\{S(n)\}$ .

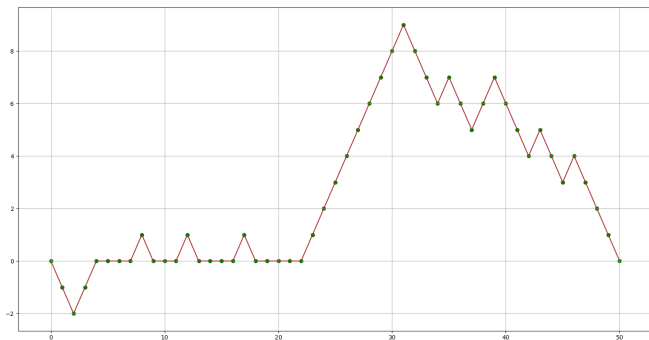
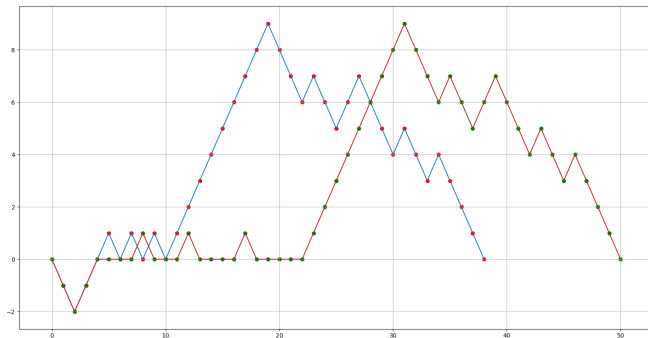


Figure:  $\tilde{S}(n)$

Two random walks together.



## The Donsker Theorem

In  $C([0, T])$

$$\frac{S(nt)}{\sqrt{n}} \xrightarrow{w} W(t), n \rightarrow \infty. \quad (1)$$

Our goal is to analyse the behaviour of a limiting random process for:

$$\tilde{X}_n(t) = \frac{\tilde{S}(nt)}{\sqrt{n}}.$$

Let times of lagging in zero be a sequence of i.i.d.r.v.  $\{\eta_n\}_{n=1}^{\infty}$  and

$$\alpha(t) = t + \sum_{i=1}^{\tau_0(nt)} \eta_i, \quad t \geq 0, \quad (2)$$

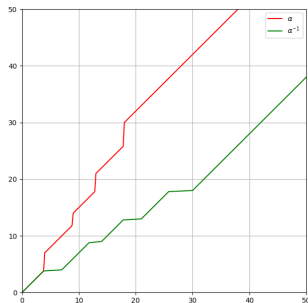
where  $\tau_0(t) = \#\{k : S(k) = 0, 0 < k \leq t\}$ , i.e. number of returns to zero before the time  $t$ .

Consider

$$\alpha^{-1}(t) = \inf\{x : \alpha(x) \geq t\}, \quad t \geq 0. \quad (3)$$

Lemma

$$\tilde{S}(t) = S(\alpha^{-1}(t)) \quad (4)$$



## Theorem 1.

Let  $\eta_i \geq 0$  be i.i.d.r.v. with  $\mathbb{E}\eta_i < \infty$ ,  $i \geq 1$ . Then the sequence  $\tilde{X}_n(t) = \frac{\tilde{S}(nt)}{\sqrt{n}}$ ,  $n \geq 1$  converges weakly in  $C([0, T])$  to a Wiener's process  $W(t)$ :

$$\tilde{X}_n(t) \xrightarrow{w} W(t), \quad n \rightarrow \infty. \quad (5)$$

Let

$$h_n(t) = \frac{\alpha^{-1}(nt)}{n},$$

then

$$\tilde{X}_n(t) = \frac{\tilde{S}(nt)}{\sqrt{n}} = \frac{S(\alpha^{-1}(nt))}{\sqrt{n}} = \frac{S(n \frac{\alpha^{-1}(nt)}{n})}{\sqrt{n}} = X_n(h_n(t)). \quad (6)$$

Recall that

$$h_n(t) = \frac{\alpha^{-1}(nt)}{n}, \quad \frac{\alpha(nt)}{n} = t + \frac{1}{n} \sum_{i=1}^{\tau_0(nt)} \eta_i, \quad \tilde{X}_n(t) = X_n(h_n(t)).$$

We will proceed by the following steps:

- 1 We will show that

$$\forall T > 0 \quad h_n(t) \stackrel{[0,T]}{\Rightarrow} t, n \rightarrow \infty, \text{ a.s.}$$

by means of showing:

- 1  $\frac{\alpha(n)}{n} \rightarrow 1, n \rightarrow \infty, \text{ a.s.}$
- 2  $\forall T > 0 \quad \frac{\alpha(nt)}{n} \stackrel{[0,T]}{\Rightarrow} t, n \rightarrow \infty, \text{ a.s.}$
- 2 Use Skorohod's representation theorem for the pair  $(X_n, h_n)$  to obtain a.s. convergence of copies, that are subjected to the same distributions:

$$X_n(t) \stackrel{n \rightarrow \infty}{\Rightarrow} W(t) \text{ a.s.}, \quad h_n(t) \stackrel{n \rightarrow \infty}{\Rightarrow} t \text{ a.s.}$$

- 3 Derive the claim on the new probability space and thus on the original one.



Consider the Markov chain  $S^p(n)$  which if it is in state 0 has probability  $p$  of leaving zero during the current step. That is

$$S^p(n) = \begin{cases} \begin{cases} +1, \text{w. pr. } 0.5 \\ S_{n-1}^p - 1, \text{w. pr. } 0.5 \end{cases} & \text{if } S_{n-1}^p \neq 0 \\ \begin{cases} +1, \text{w. pr. } 0.5 \\ -1, \text{w. pr. } 0.5 \end{cases} & \text{w. pr. } p \text{ if } S^p(n-1) = 0, \\ 0, & \text{w. pr. } 1-p \text{ if } S^p(n-1) = 0, \end{cases} \quad (7)$$

$S^p(n)$  is equivalent to  $\tilde{S}(n)$  with  $\eta_i$  geometrically distributed with parameter  $p$ . So in  $C([0, T])$

$$\frac{S^p(nt)}{\sqrt{n}} \xrightarrow{w} W(t), n \rightarrow \infty.$$

Let  $S^p(n)$  be a Markov chain on  $\mathbb{Z}$  with

$$p_{0,0} = 1 - p, \quad p_{0,\pm 1} = \frac{p}{2},$$

$$\forall i \neq 0 \quad p_{i,i\pm 1} = \frac{1}{2}.$$

And let  $p = p_n = \frac{\rho}{n^\gamma}$ . Consider

$$X_n^{p_n}(t) = \frac{S^{p_n}(nt)}{\sqrt{n}}.$$

## Theorem 2

In  $C([0, T])$ :

if  $\gamma < 0.5$ , then  $X_n^{p_n}(t) \xrightarrow[n \rightarrow \infty]{w} W(t)$ ,

if  $\gamma > 0.5$ , then  $X_n^{p_n}(t) \xrightarrow[n \rightarrow \infty]{w} 0$ ,

if  $\gamma = 0.5$ , then  $X_n^{p_n}(t) \xrightarrow[n \rightarrow \infty]{w} W_{\text{sticky}}(t)$ .

As we have seen in example,  $S^{p_n}(n) \stackrel{d}{=} \tilde{S}(n)$  with  $\eta_i^{(n)} \sim \text{Geom}(p_n)$ . So we will prove the theorem for such  $\tilde{S}(n)$ .

Now  $\alpha(t) = \alpha_n(t)$  depends on  $n$ . Still we write

$$X_n^{P_n}(t) = X_n(h_n(t)), \quad h_n(t) = \frac{\alpha_n^{-1}(nt)}{n} \quad (8)$$

and

$$\frac{\alpha_n(nt)}{n} = t + \frac{1}{n} \sum_{i=1}^{\tau_0(nt)} \eta_i^{(n)} = t + \frac{n^\gamma}{\sqrt{n}} \frac{1}{\sqrt{n}} \sum_{i=1}^{\tau_0(nt)} \frac{\eta_i^{(n)}}{n^\gamma}. \quad (9)$$

## Theorem

Let  $W(t)$  — be a Brownian motion in  $\mathbb{R}$  and  $L_W^0(t)$  — be its local time at 0, i.e

$$L_W^0(t) = \lim_{\epsilon \rightarrow 0} \frac{1}{2\epsilon} \int_0^t 1_{|W(s)| \leq \epsilon} ds.$$

Then in  $C([0, T])$ :

$$\left( \frac{\tau_0(nt)}{\sqrt{n}}, \frac{S(nt)}{\sqrt{n}} \right) \xrightarrow{w} (L_W^0(t), W(t)), \quad n \rightarrow \infty.$$

We will proceed by the following steps:

- 1 Use Skorohod's representation theorem to the pair  $(\tau_0(t), S(t))$ .
- 2 On a new probability space prove convergence of

$$\forall T > 0 \sup_{t \in [0, T]} \left| \frac{1}{\sqrt{n}} \sum_{i=1}^{\sqrt{n}L_W^0(t)} \frac{\eta_i^{(n)}}{n^\gamma} - \frac{L_W^0(t)}{\rho} \right| \xrightarrow{\mathbb{P}} 0, n \rightarrow \infty, \quad (10)$$

1

$$\frac{1}{\sqrt{n}} \sum_{i=1}^{\sqrt{nt}} \frac{\eta_i^{(n)}}{n^\gamma} \xrightarrow{\mathbb{P}} \frac{t}{\rho}, n \rightarrow \infty.$$

2

$$\forall T > 0 \sup_{t \in [0, T]} \left| \frac{1}{\sqrt{n}} \sum_{i=1}^{\sqrt{nt}} \frac{\eta_i^{(n)}}{n^\gamma} - \frac{t}{\rho} \right| \xrightarrow{\mathbb{P}} 0, n \rightarrow \infty,$$

Recall that

$$\tilde{X}_n(t) = X_n(h_n(t)), \quad \frac{\alpha_n(nt)}{n} = t + \frac{n^\gamma}{\sqrt{n}} \frac{1}{\sqrt{n}} \sum_{i=1}^{\tau_0(nt)} \frac{\eta_i^{(n)}}{n^\gamma}.$$

- ❶ For  $\gamma < 0.5$ :

$$\frac{\alpha_n(nt)}{n} \xrightarrow[n \rightarrow \infty]{w} t. \quad (11)$$

- ❷ Invoke Skorohod's theorem once again for the triplet of random elements

$$(\tau_0(t), S(t), \alpha_n(t)).$$

- ❸ Appeal to the proof of theorem 1.

$$X_n^{p_n}(t) \xrightarrow[n \rightarrow \infty]{w} W(t).$$

Recall that

$$\tilde{X}_n(t) = X_n(h_n(t)), \quad \frac{\alpha_n(nt)}{n} = t + \frac{n^\gamma}{\sqrt{n}} \frac{1}{\sqrt{n}} \sum_{i=1}^{\tau_0(nt)} \frac{\eta_i^{(n)}}{n^\gamma}.$$

- 1 For  $\gamma > 0.5$ :

for every  $\delta > 0$  on  $[\delta, T]$

$$\frac{\alpha_n(nt)}{n} \xrightarrow[n \rightarrow \infty]{w} \infty. \quad (12)$$

- 2 Invoke Skorohod's theorem once again for the triplet of random elements

$$(\tau_0(t), S(t), \alpha_n(t)).$$

- 3 Show that

$$h_n(t) \Rightarrow 0,$$

- 4 Hence the claim on the new probability space.  
5 Hence on the original one.

$$X_n^{P_n}(t) \xrightarrow[n \rightarrow \infty]{w} 0.$$

Recall that

$$\frac{\alpha_n(nt)}{n} = t + \frac{n^\gamma}{\sqrt{n}} \frac{1}{\sqrt{n}} \sum_{i=1}^{\tau_0(nt)} \frac{\eta_i^{(n)}}{n^\gamma}.$$

- ❶ For  $\gamma = 0.5$ :

$$\frac{\alpha(nt)}{n} \xrightarrow[n \rightarrow \infty]{w} t + L_W^0(t)/\rho. \quad (13)$$

- ❷ As  $\frac{\alpha(nt)}{n}$  is strictly monotonous, its generalized inverse is continuous.  
 ❸ Hence

$$\frac{\alpha^{-1}(nx)}{n} = \text{Inv}\left[\frac{\alpha(nt)}{n}\right](x) \xrightarrow{w} \text{Inv}[t + L_W^0(t)/\rho](x). \quad (14)$$

- ❹ Set

$$W_{\text{sticky}}(x) = W(\text{Inv}[t + L_W^0(t)/\rho](x))$$

*Thank you for your attention!*