

Raking-Ratio empirical process

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Introduction

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- $\alpha_n(f) = \sqrt{n}(\mathbb{P}_n(f) P(f))$ the empirical process indexed by \mathcal{F} .

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 the empirical process indexed by \mathcal{F} .

If \mathcal{F} is a Donsker class, $\alpha_n \xrightarrow[n \to +\infty]{\mathcal{L}} \mathbb{G}$ in $l^{\infty}(\mathcal{F})$ where \mathbb{G} is the *P*-brownian bridge, *i.e* the Gaussian process with covariance

$$\operatorname{Cov}(\mathbb{G}(f),\mathbb{G}(g)) = P(fg) - P(f)P(g).$$

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2. Extension 1: auxiliary information learning

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- 2. Extension 1: auxiliary information learning
- 3. Extension 2: re-sampling method with auxiliary information

Raking-ratio method

Literature: Deming/Stephan, Sinkhorn, Ireland/Kullback. **Description:**

	$A_1^{(2)}$	$A_2^{(2)}$	$A_3^{(2)}$	$\mathbb{P}_n[\mathcal{A}^{(1)}]$	$P[\mathcal{A}^{(1)}]$
$A_1^{(1)}$	0.2	0.25	0.1	0.55	0.52
$A_{2}^{(1)}$	0.1	0.2	0.15	0.45	0.48
$\mathbb{P}_n[\mathcal{A}^{(2)}]$	0.3	0.45	0.25	1	
$P[\mathcal{A}^{(1)}]$	0.31	0.4	0.29		

We have a table of frequencies whose margins do not correspond to known margins. The algorithm proposes to correct this

	A ₁ ⁽²⁾	A ₂ ⁽²⁾	A ₃ ⁽²⁾	$\mathbb{P}_n^{(1)}[\mathcal{A}^{(1)}]$	$P[\mathcal{A}^{(1)}]$
A ₁ ⁽¹⁾	0.189	0.236	0.095	0.52	0.52
A ₂ ⁽¹⁾	0.11	0.21	0.16	0.48	0.48
$\mathbb{P}_n^{(1)}[\mathcal{A}^{(2)}]$	0.299	0.446	0.255	1	
$P[\mathcal{A}^{(2)}]$	0.31	0.4	0.29		

The totals for each line are first corrected by applying a rule of three. Each cell is multiplied by the ratio of the expected total of each line on the total of each line.

	A ₁ ⁽²⁾	$A_2^{(2)}$	$A_3^{(2)}$	$\mathbb{P}_n^{(2)}[\mathcal{A}^{(1)}]$	$P[\mathcal{A}^{(1)}]$
A ₁ ⁽¹⁾	0.196	0.212	0.108	0.516	0.52
A ₂ ⁽¹⁾	0.114	0.188	0.182	0.484	0.48
$\mathbb{P}_n^{(2)}[\mathcal{A}^{(2)}]$	0.31	0.4	0.29	1	
$P[\mathcal{A}^{(2)}]$	0.31	0.4	0.29		

The same reasoning is applied to correct the totals for each column. These last two operations are repeated in a loop.

	A ₁ ⁽²⁾	A ₂ ⁽²⁾	A ₃ ⁽²⁾	$\mathbb{P}_n^{(\infty)}[\mathcal{A}^{(1)}]$	$P[\mathcal{A}^{(1)}]$
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A ₂ ⁽¹⁾	0.111	0.188	0.181	0.48	0.48
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$P[\mathcal{A}^{(2)}]$	0.31	0.4	0.29		

Very quickly the algorithm stabilizes. Totals are the expected totals. For this example it took only 7 iterations.

Remark: we can rake on more than two partitions!

In turn *N* the algorithm does:

$$p^{(N+1)}(A) = \sum_{j=1}^{m_{N+1}} p^{(N)}(A \cap A_j^{(N+1)}) \frac{P(A_j^{(N+1)})}{p^{(N)}(A_j^{(N+1)})}.$$

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We define the raked empirical measure $\mathbb{P}_n^{(N)}$ to be $\mathbb{P}_n^{(0)} = \mathbb{P}_n$ and

$$\mathbb{P}_{n}^{(N+1)}(f) = \sum_{j=1}^{m_{N+1}} \mathbb{P}_{n}^{(N)}(f\mathbf{1}_{A_{j}^{(N+1)}}) \frac{P(A_{j}^{(N+1)})}{\mathbb{P}_{n}^{(N)}(A_{j}^{(N+1)})}$$

In particular, $\mathbb{P}_n^{(N+1)}(A_j^{(N+1)}) = P(A_j^{(N+1)}), \forall j = 1, \dots, m_{N+1}.$

Let $\alpha_n^{(N)}(f) = \sqrt{n}(\mathbb{P}_n^{(N)}(f) - P(f))$ the raked empirical process.

$$\alpha_n^{(N+1)}(f) = \sum_{j \leqslant m_{N+1}} \frac{P(A_j^{(N+1)})}{\mathbb{P}_n^{(N)}(A_j^{(N+1)})} \left(\alpha_n^{(N)}(f\mathbf{1}_{A_j^{(N+1)}}) - \mathbb{E}[f|A_j^{(N+1)}]\alpha_n^{(N)}(A_j^{(N+1)}) \right)$$

with $\mathbb{E}[f|A] = \frac{P(f\mathbf{1}_A)}{P(A)}$.

In particular, $\alpha_n^{(N+1)}(A_j^{(N+1)}) = 0$, $\forall j = 1, ..., m_{N+1}$.

Remark: $\alpha_n^{(N)}$ is no more centered.

- Properties of $\alpha_n^{(N)}(\mathcal{F})$;

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- Weak convergence in $\ell^{\infty}(\mathcal{F})$ of $\alpha_n^{(N)}(\mathcal{F})$ when $n \to +\infty$ towards a centered Gaussian process $\mathbb{G}^{(N)}(\mathcal{F})$;

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- Variance of G^(N)(f): is it lower than that of G? If a loop is performed with the Raking-Ratio method, does the variance decrease with each loop turn?

- Properties of $\alpha_n^{(N)}(\mathcal{F})$;
- Weak convergence in $\ell^{\infty}(\mathcal{F})$ of $\alpha_n^{(N)}(\mathcal{F})$ when $n \to +\infty$ towards a centered Gaussian process $\mathbb{G}^{(N)}(\mathcal{F})$;
- Variance of $\mathbb{G}^{(N)}(f)$: is it lower than that of \mathbb{G} ? If a loop is performed with the Raking-Ratio method, does the variance decrease with each loop turn?
- If we rake only two partitions, what's the limit of $\alpha_n^{(N)}(\mathcal{F})$ as $n, N \to +\infty$?

Law of iterated logarithm

If \mathcal{F} satisfies some entropy conditions then for all $N_0 \in \mathbb{N}$,

$$\limsup_{n \to +\infty} \sqrt{\frac{n}{LLn}} \max_{0 \leq N \leq N_0} ||\mathbb{P}_n^{(N)} - P||_{\mathcal{F}} \leq \sqrt{2}\sigma_{\mathcal{F}} \prod_{N=1}^{N_0} \left(1 + \frac{M}{\delta_N}\right) \text{a.s.},$$

where

- $\delta_N = \min_{j \leq m_N} P(A_j^{(N)})$;
- $\sigma_{\mathcal{F}}^2 = \sup_{\mathcal{F}} \operatorname{Var}(f)$;
- $M = ||f||_{\mathcal{F}}$.

Recall that

$$\limsup_{n \to +\infty} \sqrt{\frac{n}{LLn}} ||\mathbb{P}_n - P||_{\mathcal{F}} \leq \sqrt{2}\sigma_{\mathcal{F}}.$$

Talagrand inequality

If \mathcal{F} satisfies some entropy conditions then for all $N_0 \in \mathbb{N}$ and $t > t_0$,

$$\mathbb{P}\left(\max_{0 \leqslant N \leqslant N_0} ||\alpha_n^{(N)}||_{\mathcal{F}} > t\right) \leqslant D_1 \exp(-D_2 t^2),$$

or

$$\mathbb{P}\left(\max_{0\leqslant N\leqslant N_0}||\alpha_n^{(N)}||_{\mathcal{F}}>t\right)\leqslant D_1t^{\nu}\exp(-D_2t^2),$$

for some $D_1, D_2, \nu > 0$.

Weak convergence of $\alpha_n^{(N)}$

Under some entropy conditions on \mathcal{F} ,

$$(\alpha_n^{(0)},\ldots,\alpha_n^{(N_0)}) \xrightarrow[n \to +\infty]{\mathcal{L}} (\mathbb{G}^{(0)},\ldots,\mathbb{G}^{(N_0)}) \quad \text{in} \quad \ell^{\infty}(\mathcal{F}^{N_0} \to \mathbb{R}^{N_0})$$

with $\mathbb{G}^{(N)}$ the Gaussian process defined by

$$\mathbb{G}^{(0)} = \mathbb{G}$$
 and $\mathbb{G}^{(N+1)}(f) = \mathbb{G}^{(N)}(f) - \sum_{j=1}^{m_{N+1}} \mathbb{E}[f|A_j^{(N+1)}]\mathbb{G}^{(N)}(A_j^{(N+1)})$

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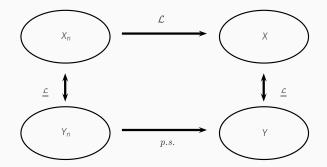
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Recall that

$$\alpha_n^{(N+1)}(f) = \sum_{j \le m_{N+1}} \frac{P(A_j^{(N+1)})}{\mathbb{P}_n^{(N)}(A_j^{(N+1)})} \left(\alpha_n^{(N)}(f_{A_j^{(N+1)}}) - \mathbb{E}[f|A_j^{(N+1)}]\alpha_n^{(N)}(A_j^{(N+1)})\right)$$

Spirit of strong approximation



Results: KMT, Berthet-Mason.

Strong approximation of $\alpha_n^{(N)}(\mathcal{F})$

Under some entropy conditions on \mathcal{F} we can construct on the same probability space X_1, \ldots, X_n and a version $\mathbb{G}_n^{(N)}$ of $\mathbb{G}^{(N)}$ such that for large n,

$$\mathbb{P}\left(\max_{0\leqslant N\leqslant N_{0}}||\alpha_{n}^{(N)}-\mathbb{G}_{n}^{(N)}||_{\mathcal{F}}>Cv_{n}\right)\leqslant\frac{1}{n^{2}},$$

with $v_n \rightarrow 0$.

By Borell-Cantelli,

$$\max_{0 \leq N \leq N_0} ||\alpha_n^{(N)} - \mathbb{G}^{(N)}||_{\mathcal{F}} = O_{\text{p.s.}}(V_n).$$

Berry-Esseen bound

Under some entropy conditions on \mathcal{F} ,

$$\max_{0 \le N \le N_0} \sup_{f \in \mathcal{F}} \sup_{x \in \mathbb{R}} \left| \mathbb{P}(\alpha_n^{(N)}(f) \le x) - \mathbb{P}(\mathbb{G}^{(N)}(f) \le x) \right| \le Cv_n.$$

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Bias and variance estimation

Under some entropy conditions on \mathcal{F} , there exists C > 0 such that

$$\limsup_{n \to +\infty} \frac{\sqrt{n}}{v_n} \max_{0 \le N \le N_0} \sup_{f \in \mathcal{F}} \left| \mathbb{E}[\mathbb{P}_n^{(N)}(f)] - P(f) \right| \le C,$$
$$\limsup_{n \to +\infty} \frac{n}{v_n} \sup_{f \in \mathcal{F}} \left| \operatorname{Var}(\mathbb{P}_n^{(N)}(f)) - \frac{1}{n} \operatorname{Var}(\mathbb{G}^{(N)}(f)) \right| \le C.$$

Raking-Ratio results

We denote

•
$$\mathbb{E}[f|\mathcal{A}^{(k)}] = (\mathbb{E}[f|\mathcal{A}^{(k)}_{1}], \dots, \mathbb{E}[f|\mathcal{A}^{(k)}_{m_k}])^t$$
;

•
$$\mathbb{G}[\mathcal{A}^{(k)}] = (\mathbb{G}(\mathcal{A}_1^{(k)}), \dots, \mathbb{G}(\mathcal{A}_{m_k}^{(k)}))^t;$$

•
$$(\mathsf{P}_{\mathcal{A}^{(k)}|\mathcal{A}^{(l)}})_{i,j} = P(\mathsf{A}_j^{(k)}|\mathsf{A}_i^{(l)}).$$

Expression of $\mathbb{G}^{(\textit{N})}$

For all $N \in \mathbb{N}^*$ and $f \in \mathcal{F}$ it holds

$$\mathbb{G}^{(N)}(f) = \mathbb{G}(f) - \sum_{k=1}^{N} \Phi_k^{(N)}(f)^t \cdot \mathbb{G}[\mathcal{A}^{(k)}]$$

where

$$\Phi_{k}^{(N)}(f) = \mathbb{E}[f|\mathcal{A}^{(k)}] + \sum_{\substack{1 \leq l \leq N-k \\ k < l_{1} < \cdots < l_{L} \leq N}} (-1)^{L} \mathsf{P}_{\mathcal{A}^{(l_{1})}|\mathcal{A}^{(k)}} \mathsf{P}_{\mathcal{A}^{(l_{2})}|\mathcal{A}^{(l_{1})}} \cdots \mathsf{P}_{\mathcal{A}^{(l_{L})}|\mathcal{A}^{(l_{L}-1)}} \cdot \mathbb{E}[f|\mathcal{A}^{(l_{L})}].$$

We denote $(\operatorname{Var}((X_1,\ldots,X_n)^t))_{i,j} = \operatorname{Cov}(X_i,X_j)$

Variance and covariance of $\mathbb{G}^{(N)}$ For all $N \in \mathbb{N}^*$ and $f, g \in \mathcal{F}$ it holds

$$\begin{aligned} \operatorname{Var}(\mathbb{G}^{(N)}(f)) &= \operatorname{Var}(\mathbb{G}(f)) - \sum_{k=1}^{N} \Phi_{k}^{(N)}(f)^{t} \cdot \operatorname{Var}(\mathbb{G}[\mathcal{A}^{(k)}]) \cdot \Phi_{k}^{(N)}(f) \\ \operatorname{Cov}(\mathbb{G}^{(N)}(f), \mathbb{G}^{(N)}(g)) &= \operatorname{Cov}(\mathbb{G}(f), \mathbb{G}(g)) \\ &- \sum_{k=1}^{N} \operatorname{Cov}\left(\Phi_{k}^{(N)}(f)^{t} \cdot \mathbb{G}[\mathcal{A}^{(k)}], \Phi_{k}^{(N)}(g)^{t} \cdot \mathbb{G}[\mathcal{A}^{(k)}]\right) \end{aligned}$$

Raking-Ratio results

Corollary 1

For any $N \in \mathbb{N}$ and $f \in \mathcal{F}$, $Var(\mathbb{G}^{(N)}(f)) \leq Var(\mathbb{G}(f))$.

For any $\{f_1, \ldots, f_m\} \in \mathcal{F}, \Sigma_m - \Sigma_m^{(N)}$ is positive definite with

$$\begin{split} \boldsymbol{\Sigma}_n^{(N)} &= \operatorname{Var}((\mathbb{G}^{(N)}(f_1), \dots, \mathbb{G}^{(N)}(f_m))^t), \\ \boldsymbol{\Sigma}_n &= \operatorname{Var}((\mathbb{G}(f_1), \dots, \mathbb{G}(f_m))^t). \end{split}$$

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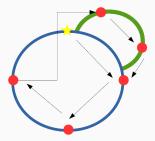
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Corollary 2 Let $N_0, N_1 \in \mathbb{N}$ s.t. $N_1 \ge 2N_0$ and $\mathcal{A}^{(N_0-i)} = \mathcal{A}^{(N_1-i)}, \quad \forall 0 \le i < N_0.$ Then for all $f \in \mathcal{F}$, $\operatorname{Var}(\mathbb{G}^{(N_1)}(f)) \le \operatorname{Var}(\mathbb{G}^{(N_0)}(f)).$



Results for 2 margins

We note $\mathcal{A} = \mathcal{A}^{(2)} = \{A_1, \dots, A_{m_1}\}$ and $\mathcal{B} = \mathcal{A}^{(1)} = \{B_1, \dots, B_{m_2}\}.$

Expression of $\mathbb{G}^{(N)}$

Let $N \in \mathbb{N}$ and $f \in \mathcal{F}$. Then for $m \in \mathbb{N}$,

$$\mathbb{G}^{(2m)}(f) = \mathbb{G}(f) - \left(S_{1,even}^{(m-1)}(f)\right)^{t} \mathbb{G}[\mathcal{A}] - \left(S_{2,even}^{(m-2)}(f)\right)^{t} \mathbb{G}[\mathcal{B}]$$
$$\mathbb{G}^{(2m+1)}(f) = \mathbb{G}(f) - \left(S_{1,odd}^{(m-1)}(f)\right)^{t} \mathbb{G}[\mathcal{A}] - \left(S_{2,odd}^{(m-1)}(f)\right)^{t} \mathbb{G}[\mathcal{B}]$$

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Recall that

$$\mathbb{G}^{(N)}(f) = \mathbb{G}(f) - \sum_{k=1}^{N} \Phi_k^{(N)}(f)^t \cdot \mathbb{G}[\mathcal{A}^{(k)}]$$

Hypothesis $\text{Matrices } \mathsf{P}_{\mathcal{A}|\mathcal{B}}\mathsf{P}_{\mathcal{B}|\mathcal{A}} \text{ and } \mathsf{P}_{\mathcal{B}|\mathcal{A}}\mathsf{P}_{\mathcal{A}|\mathcal{B}} \text{ are ergodic.}$

Hypothesis

Matrices $P_{\mathcal{A}|\mathcal{B}}P_{\mathcal{B}|\mathcal{A}}$ and $P_{\mathcal{B}|\mathcal{A}}P_{\mathcal{A}|\mathcal{B}}$ are ergodic.

Convergence of $S_{i,even}^{(N)}(f), S_{i,odd}^{(N)}(f)$

 $S_{i,even}^{(N)}(f), S_{i,odd}^{(N)}(f)$ for i = 1, 2 converge respectively towards $S_{i,even}(f), S_{i,odd}(f)$. They verify the relations:

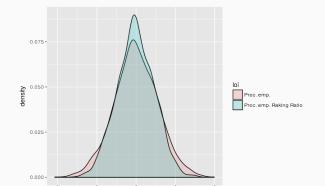
$$S_{1,odd}(f) = S_{1,even}(f) + \begin{pmatrix} \mathbb{E}[f] \\ \vdots \\ \mathbb{E}[f] \end{pmatrix}, \qquad S_{2,even}(f) = S_{2,odd}(f) + \begin{pmatrix} \mathbb{E}[f] \\ \vdots \\ \mathbb{E}[f] \end{pmatrix}.$$

Results for 2 margins

Convergence of $\mathbb{G}^{(N)}$

The sequence of process $(\mathbb{G}^{(N)})_N$ converges in distribution when $N \to +\infty$ to the centered Gaussian process $\mathbb{G}^{(\infty)}$ indexed by \mathcal{F} and defined by

$$\mathbb{G}^{(\infty)}(f) = \mathbb{G}(f) - S_{1,even}(f)^{t} \cdot \mathbb{G}[\mathcal{A}] - S_{2,odd}(f)^{t} \cdot \mathbb{G}[\mathcal{B}]$$



Extension 1: auxiliary information learning

Motivation

We suppose that the auxiliary information is given by an estimate of the probability of belonging to a set of several partitions.

The auxiliary information is given by

 $\mathbb{P}'_{N}[\mathcal{A}^{(N)}] = (\mathbb{P}_{n}(A_{1}^{(N)}), \dots, \mathbb{P}'_{N}(A_{m_{N}}^{(N)})),$

a multinomial distribution with $n_N > 0$ trials and event probabilities

$$P[\mathcal{A}^{(N)}] = (P(A_1^{(N)}), \dots, P(A_{m_N}^{(N)})).$$

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$$P[\mathcal{A}^{(N)}] = (P(A_1^{(N)}), \dots, P(A_{m_N}^{(N)})).$$

Goal

We study the raking-ratio empirical process which uses $\mathbb{P}'_{N}[\mathcal{A}^{(N)}]$ instead of $P[\mathcal{A}^{(N)}]$.

Definition of $\widetilde{\mathbb{P}}_n^{(N)}(\mathcal{F})$

We define the *N*-th raking-ratio empirical measure with auxiliary information learning $\widetilde{\mathbb{P}}_{n}^{(N)}(\mathcal{F})$ as the same way as $\mathbb{P}_{n}^{(N)}(\mathcal{F})$: $\widetilde{\mathbb{P}}_{n}^{(0)}(f) = \mathbb{P}_{n}(f)$ and for $N \ge 1$,

$$\widetilde{\mathbb{P}}_{n}^{(N)}(f) = \sum_{j=1}^{m_{N}} \frac{\mathbb{P}_{N}'(A_{j}^{(N)})}{\widetilde{\mathbb{P}}_{n}^{(N-1)}(A_{j}^{(N)})} \widetilde{\mathbb{P}}_{n}^{(N-1)}(f|_{A_{j}^{(N)}}).$$

Notice that

$$\widetilde{\mathbb{P}}_n[\mathcal{A}^{(N)}] = \left(\widetilde{\mathbb{P}}_n^{(N)}(\mathcal{A}_1^{(N)}), \dots, \widetilde{\mathbb{P}}_n^{(N)}(\mathcal{A}_{m_N}^{(N)})\right) = \mathbb{P}'_N[\mathcal{A}^{(N)}].$$

Recall that

$$\mathbb{P}_{n}^{(N)}(f) = \sum_{j=1}^{m_{N}} \frac{\mathbb{P}_{n}(A_{j}^{(N)})}{\mathbb{P}_{n}^{(N-1)}(A_{j}^{(N)})} \mathbb{P}_{n}^{(N-1)}(f_{A_{j}^{(N)}}).$$

Definition of $\widetilde{\alpha}_n^{(N)}(\mathcal{F})$

We define the *N*-th raking-ratio empirical process with estimated auxiliary information by

$$\widetilde{\alpha}_n^{(N)}(f) = \sqrt{n}(\widetilde{\mathbb{P}}_n^{(N)}(f) - P(f)).$$

Notice that $\alpha_n^{(N)}(A_j^{(N)}) \neq 0$.

Strong approximation of $\alpha_n^{(N)}(\mathcal{F})$

Under some entropy conditions on \mathcal{F} we can construct on the same probability space X_1, \ldots, X_n and a version $\mathbb{G}_n^{(N)}$ of $\mathbb{G}^{(N)}$ such that for large n,

$$\mathbb{P}\left(\max_{0\leqslant N\leqslant N_0}||\widetilde{\alpha}_n^{(N)}-\mathbb{G}_n^{(N)}||_{\mathcal{F}}>C\left(V_n+\sqrt{\frac{n\log(n)}{n_{(N_0)}}}\right)\right)\leqslant \frac{1}{n^2}$$

with $v_n \rightarrow 0$ and $n_{(N_0)} = \min_{N \leq N_0} n_N$.

Extension 2: re-sampling method with auxiliary information

Notation

Bootstrap is a statistical method for re-sampling. It replaces P by \mathbb{P}_n .

A general way to define the bootstrap is to multiply $f(X_i)$ by a random variable Z_i such that $\mathbb{E}[Z_i|X_i] = 1$ and $\operatorname{Var}(Z_i) = 1$.

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A general way to define the bootstrap is to multiply $f(X_i)$ by a random variable Z_i such that $\mathbb{E}[Z_i|X_i] = 1$ and $\operatorname{Var}(Z_i) = 1$.

We define the bootstrapped empirical measure and process:

$$\mathbb{P}_n^*(f) = \frac{1}{\sum_{i=1}^n Z_i} \sum_{i=1}^n Z_i f(X_i), \qquad \alpha_n^*(f) = \sqrt{n} (\mathbb{P}_n^*(f) - \mathbb{P}_n(f)).$$

Notation

Bootstrap is a statistical method for re-sampling. It replaces P by \mathbb{P}_n .

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Goal

- Make the strong approximation of α_n^* to $\mathbb{G}^*,$ a $P\text{-}\mathsf{Brownian}$ bridge independent of \mathbb{G} ;
- Bootstrap the Raking-Ratio empirical process to simulate its distribution.

Strong approximation of α_n^*

Under some entropy conditions on \mathcal{F} we can construct on the same probability space (X_n, Z_n) and $(\mathbb{G}_{n, *} \mathbb{G}_n^*)$ of *P*-Brownian bridge such that for large *n*,

$$\mathbb{P}\left(\{||\alpha_n - \mathbb{G}_n||_{\mathcal{F}} > C \mathsf{v}_n\} \bigcup \{||\alpha_n^* - \mathbb{G}_n^*||_{\mathcal{F}} > C \mathsf{v}_n\}\right) \leq \frac{1}{n^2},$$

with $v_n \rightarrow 0$ depends on the entropy of (\mathcal{F}, P) .

Goal

How can we adapt the bootstrap method to simulate the distribution of the Raking-Ratio empirical process?

 $\mathbb{P}_n^{*(0)} = \mathbb{P}_n^*$ and

$$\mathbb{P}_{n}^{*(N+1)}(f) = \sum_{j=1}^{m_{N+1}} \mathbb{P}_{n}^{*(N)}(f\mathbf{1}_{A_{j}^{(N+1)}}) \frac{\mathbb{P}_{n}(A_{j}^{(N+1)})}{\mathbb{P}_{n}^{*(N)}(A_{j}^{(N+1)})},$$

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Result

 $\alpha_n^{*(N)} \to \mathbb{G}^{*(N)}$ in $\ell^{\infty}(\mathcal{F})$ and $\mathbb{G}^{*(N)}$ has the same distribution as $\mathbb{G}^{(N)}$.

Thank you for your attention!

Questions?